

Research Article

Modeling Temperature and Pricing Weather Derivatives Based on Temperature

Birhan Taştan¹ and Azize Hayfavi²

¹*Department of Banking and Finance, Faculty of Management, Girne American University, 99428 Karmi Campus, University Drive, P.O. Box 5, Karaoglanoglu, Kyrenia, Northern Cyprus, Mersin 10, Turkey*

²*Department of Financial Mathematics, Institute of Applied Mathematics, Middle East Technical University, Üniversiteler Mahallesi, Dumlupınar Bulvarı No. 1, Çankaya, 06800 Ankara, Turkey*

Correspondence should be addressed to Birhan Taştan; birhantastan@gau.edu.tr

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This study first proposes a temperature model to calculate the temperature indices upon which temperature-based derivatives are written. The model is designed as a mean-reverting process driven by a Levy process to represent jumps and other features of temperature. Temperature indices are mainly measured as deviations from a base temperature, and, hence, the proposed model includes jumps because they may constitute an important part of this deviation for some locations. The estimated value of a temperature index and its distribution in this model apply an inversion formula to the temperature model. Second, this study develops a pricing process over calculated index values, which returns a customized price for temperature-based derivatives considering that temperature has unique effects on every economic entity. This personalized price is also used to reveal the trading behavior of a hypothesized entity in a temperature-based derivative trade with profit maximization as the objective. Thus, this study presents a new method that does not need to evaluate the risk-aversion behavior of any economic entity.

1. Introduction

Temperature-based derivatives represent a new financial tool to buy and sell a natural phenomenon. Doing so requires two things: a unit of measurement for the natural phenomenon that everyone agrees upon and a price that may facilitate a transaction. This study is designed to evaluate these two requirements.

Some preexisting measures already appear in the form of indices to meet the first requirement. To find values for these indices, the literature contains several temperature models using mean-reverting processes as the main tool. The most cited study develops an Ornstein-Uhlenbeck (OU) process to model temperature [1]. Using the equivalent martingale measures approach, the authors determine the price of an option. Benth and Šaltytė-Benth [2] model temperature as a continuous time autoregressive process for Stockholm and report a clear seasonal variation in regression residuals. They propose a model using a higher-order continuous

time autoregressive process, driven by a Wiener process with seasonal standard deviation. While pricing futures and options, they consider a Gaussian structure in the temperature dynamics. In another study [3], they model temperature with an OU process driven by a generalized Levy process. The model contains seasonal mean and volatility. Instead of dynamic models, some authors offer time-series models to represent temperature. Campbell and Diebold [4] apply a time series approach to model temperature, including trend seasonality represented by a low-ordered Fourier series and cyclical patterns represented by autoregressive lags. The contributions to conditional variance dynamics are coming from seasonal and cyclical components. The authors used Fourier series and GARCH processes to represent seasonal volatility components and cyclical volatility components, respectively. Jewson and Caballero [5] discuss the use of weather forecasts in pricing weather derivatives, presenting two methods for strong seasonality in probability distributions and the autocorrelation structure of temperature anomalies. Elias et al.

[6] develop four regime-switching models of temperature for pricing temperature based derivatives and find that a two-state model governed by a mean-reverting process as the first state and by a Brownian motion as the second state was superior to the others. Schiller et al. [7] and Oetomo and Stevenson [8] provide a comparison of different models.

To comply with the first requirement, the current study offers a temperature model based on Alaton et al. [1], which was defined after analyzing temperature data from different locations. The temperature model in this study is a mean-reverting Levy process. The Levy part contains a Brownian motion and two mean reverting jump processes driven by compound Poisson processes. For some flexibility, the jumps are designed as slow and fast mean-reverting processes, which are independent. The main difference with the model proposed here is its inclusion of jumps. Because temperature indices are mainly calculated as deviation of temperature from a base temperature, the model assumes that jumps are inevitable, at least for certain locations. The numerical estimates in this study contain test results related to this issue. The solution to the proposed temperature model is applied inversion formula to obtain approximated expected value of a specific index type and to obtain the approximated distribution of the same index.

Notably, temperature has unique behavior for any location in which it is measured. Therefore, it is not possible to develop a single model that explains every temperature behavior in every location. In addition, more than 100,000 weather stations worldwide measure temperature for different periods. It may even be difficult to develop a temperature model that is valid for all time at a single location. Thus, this study aims to cover more locations and periods by simply using a flexible model that can include or exclude jumps.

The second requirement, temperature-based derivatives pricing, is more complicated. Because the underlying commodity is not a traded asset, weather derivatives based on temperature have an incomplete market [9]. Carr et al. [10], Magill and Quinzii [11], and El Karoui and Quenez [12] provide a general discussion of incomplete markets. Pricing temperature-based derivatives is mainly based on two approaches: dynamic valuation and equilibrium asset pricing. The dynamic pricing approaches [1, 2] were discussed above. For equilibrium pricing, Cao and Wei [13] use a generalization of the Jr. Lucas model [14], which considers weather as another source of uncertainty. Richards et al. [15] suggest another equilibrium model. Davis [16] uses the marginal substitution value approach for pricing in incomplete markets. In addition, Xu et al. [17] use another classification for pricing temperature-based derivatives and add actuarial pricing and extended risk-neutral valuation in addition to equilibrium asset pricing, where the former is based on Jewson and Brix [18] and the latter on Hull [19] and Turvey [20]. In addition, some researchers used Monte-Carlo simulations in pricing temperature-based derivatives [21].

This study bases its pricing on the monetary effect of the natural phenomenon on economic entities. Further, this study shows that temperature has different effects on different entities. The same temperature may have a positive effect on one entity and a negative effect on another and is therefore

personal, *ceteris paribus*. Thus, the study develops a personal price, which may require a determination of the entity's risk aversion behavior. To address this problem, this study focuses on entity-specific trading behavior rather than the entity's risk aversion behavior in order to develop a more realistic approach by avoiding an inconclusive debate over the risk premiums and utility functions used to calculate risk premiums. Moreover, a benefit of the proposed pricing model is that it is independent from how researchers measure temperature.

Critics may object to the move from a stochastic temperature model to some form of actuarial pricing model. There are several reasons for this move: first, this study demonstrates that risk-neutral pricing ends up with superhedging; second, the discussion about risk premiums in the literature is unclear; and, finally, the calculations of jump processes needed approximations to obtain certain results. These considerations led to this study's development of a more appropriate and practical method.

The second section of the paper provides the approximated index calculation and distribution of temperature after presenting a temperature model. Third section develops individualized prices and discusses the trading behavior of a hypothetical entity. The paper then presents the study's conclusions.

2. Model

Some basic terminology is defined in the following:

$$T_i = \frac{T_i^{\max} + T_i^{\min}}{2}, \quad (1)$$

where T is the daily average temperature, i represents a certain day, and T_i^{\max} and T_i^{\min} are the maximum and minimum temperatures of the given day, respectively.

For the Heating Degree Day (HDD) temperature index used in temperature-based derivatives, for a given day,

$$\text{HDD}_i = \max(0, \text{Base} - T_i), \quad (2)$$

where Base is a predetermined temperature level and T_i is the average temperature calculated as in (1) for a given day i .

Cumulative HDD (CHDD):

$$\text{CHDD} = \sum_{i=1}^N \text{HDD}_i, \quad (3)$$

where HDD_i is calculated as in (2) and N is the time horizon, which is generally a month or a season.

2.1. The Temperature Model. Based on Alaton et al. [1], the temperature model is an OU process driven by a Levy process that contains independent processes as Brownian motions and two mean-reverting compound Poisson processes. The

model is represented as follows:

$$dT_t = \left\{ \frac{dT_t^m}{dt} + b(T_t^m - T_t) \right\} dt + dL_t, \quad (4)$$

where T_t^m is a cyclical process of temperature and represented in (5). Additionally, b is the mean-reversion parameter, and subscript t represents time.

$$T_t^m = A + Bt + C \sin(wt + \varphi) \quad (5)$$

where $w = \frac{2\pi}{365}$, φ is the phase angle.

The differential of the driving Levy process dL_t is defined as follows:

$$dL_t = \sigma_t dW_t + dY_t + dZ_t. \quad (6)$$

The Brownian component of L_t will be approximated by the ARCH (1) model. To represent the different jump structures in temperature in the form of a single jump and a series of jumps, dY_t and dZ_t are defined as fast and slow mean-reverting OU processes driven by compound Poisson processes with intensities of λ_Y and λ_Z , and α and β being mean-reversion parameters, respectively. Hayfavi and Talasli [22] use a somewhat similar mean-reverting jump process combination in their model of spot electricity prices.

$$dY_t = -\alpha Y_t dt + dQ_t \quad \text{where } Q_t = \sum_{i=1}^{N_t^Y} U_i, \quad U_i \text{ are iid random variables, } U_i \sim N(\mu_Y, \delta_Y^2), \quad (7)$$

$$dZ_t = -\beta Z_t dt + dR_t \quad \text{where } R_t = \sum_{i=1}^{N_t^Z} V_i, \quad V_i \text{ are iid random variables, } V_i \sim N(\mu_Z, \delta_Z^2).$$

The solutions to these non-Gaussian processes are [23] the following:

$$Y_t = y_0 e^{-\alpha t} + \int_0^t e^{\alpha(s-t)} dQ_s, \quad (8)$$

$$Z_t = z_0 e^{-\beta t} + \int_0^t e^{\beta(s-t)} dR_s.$$

The solution to (4) is given as

$$T_t = T_t^m + e^{-bt} (T_0 - T_0^m) + e^{-bt} \int_0^t e^{bu} dL_u. \quad (9)$$

To find the value of a temperature-based derivative, one needs the distribution of the underlying temperature given in (9). However, this does not have a closed-form solution. One way to address this problem is to use a characteristic function of the temperature and apply inversion techniques to find the value of an HDD, an approximated distribution of CHDD, and an approximated distribution for temperature itself.

2.2. Characteristic Function of Temperature. This study follows Cont and Tankov [23] to find the characteristic function given in (9). First, using L_1 , the characteristic exponent of (6) will be determined, where characteristic exponent $\psi(u)$ is defined as $Ee^{iuL} = e^{\psi(u)}$. The solution to L_1 is

$$L_1 = \int_0^1 \sigma_u dW_u + y_0 e^{-\alpha} + \int_0^1 e^{\alpha(s-1)} dQ_s + z_0 e^{-\beta} + \int_0^1 e^{\beta(s-1)} dR_s. \quad (10)$$

Then, the characteristic exponents of the Levy components will be

$$\psi_{BM}(u) = -\frac{1}{2} u^2 C, \quad (11)$$

where $C = E \left(\int_0^1 \sqrt{\sigma_u^2} dW_u \right)^2 = \int_0^1 \sigma_u^2 du$

and for jump processes

$$\psi_Y(u) = iu y_0 e^{-\alpha} + \lambda_Y \int_0^1 \left\{ e^{iu \mu_Y e^{-\alpha(r-1)} - (1/2) u^2 \delta_Y^2 e^{2\alpha(r-1)}} - 1 \right\} dr. \quad (12)$$

$\psi_Z(u)$ can be written similarly. It is not possible to evaluate the integral in (12). Consequently, the following approximation method was developed.

Let $A = iu \mu_Y e^{-\alpha}$ and $B = (1/2) u^2 \delta_Y^2 e^{-2\alpha}$. Let $e^{\alpha r} = g(r)$. Then, $\psi_Y(u) = iu y_0 e^{-\alpha} + \lambda_Y \int_0^1 \{e^{Ag(r) - Bg^2(r)} - 1\} dr$.

Further, let $D(r) = Ag(r) - Bg^2(r)$, where $e^{D(r)} = 1 + D(r) + D^2(r)/2! + \dots$.

Then, by using linear approximation,

$$\psi_Y(u) = iu y_0 e^{-\alpha} + \lambda_Y iu \mu_Y \left(\frac{1 - e^{-\alpha}}{\alpha} \right) - \lambda_Y \frac{1}{2} u^2 \delta_Y^2 \left(\frac{1 - e^{-2\alpha}}{2\alpha} \right). \quad (13)$$

Again, $\psi_Z(u)$ can be written similarly. Finally, the characteristic function of the temperature model can be written explicitly. Referring to Cont and Tankov [23],

$$E\{e^{iuT_t}\} = \exp\left\{iu\left(T_t^m + e^{-bt}(T_0 - T_0^m)\right) + \int_0^t \psi_T(ue^{b(s-t)})ds\right\}. \quad (14)$$

In explicit form,

$$\begin{aligned} E\{e^{iuT_t}\} &= \theta_T(u) = \exp\left[iu\left(T_t^m + e^{-bt}(T_0 - T_0^m)\right) + iu\left(\frac{1 - e^{-bt}}{b}\right)\left\{y_0e^{-\alpha} + \lambda_y\mu_y\left(\frac{1 - e^{-\alpha}}{\alpha}\right) + z_0e^{-\beta} + \lambda_z\mu_z\left(\frac{1 - e^{-\beta}}{\beta}\right)\right\} - \frac{1}{2}u^2\left(\frac{1 - e^{-2bt}}{2b}\right) \cdot \left\{C + \lambda_y\delta_Y^2\left(\frac{1 - e^{-2\alpha}}{2\alpha}\right) + \lambda_z\delta_z^2\left(\frac{1 - e^{-2\beta}}{2\beta}\right)\right\}\right]. \end{aligned} \quad (15)$$

2.3. HDD and Distribution Function. This part of the study focuses on measuring HDDs. It is easy to apply the calculations into other types of indices. In the current case, inversion techniques will be used to find the value of an HDD and its distribution and hence the CHDD values.

2.3.1. Approximating Density Function of Temperature. To find an approximating density function of temperature, inversion formula will be applied to the characteristic function of the temperature defined in (15). Before applying the inversion formula, the following shortcuts are derived from (15). Let $f(x)$ and $\theta(z)$ be the density function and characteristic function of temperature, respectively.

$$T^* = (T_t^m + e^{-bt}(T_0 - T_0^m)) \quad (16)$$

$$M = \left(\frac{1 - e^{-bt}}{b}\right)\left\{y_0e^{-\alpha} + \lambda_y\mu_y\left(\frac{1 - e^{-\alpha}}{\alpha}\right) + z_0e^{-\beta} + \lambda_z\mu_z\left(\frac{1 - e^{-\beta}}{\beta}\right)\right\} \quad (17)$$

$$M^* = T^* + M \quad (18)$$

$$V = \left(\frac{1 - e^{-2bt}}{2b}\right)\left\{C + \lambda_y\delta_Y^2\left(\frac{1 - e^{-2\alpha}}{2\alpha}\right) + \lambda_z\delta_z^2\left(\frac{1 - e^{-2\beta}}{2\beta}\right)\right\}. \quad (19)$$

Then, by inversion formula $f(x) = (1/2\pi) \int_{-\infty}^{\infty} e^{-izx} \theta(z) dz$, the result is

$$f(x) = \frac{1}{\sqrt{2\pi V}} e^{-(x - M^*)^2/2V}. \quad (20)$$

Because weather derivatives are defined on CHDDs, their distribution will be defined. Clearly, HDD values may show autocorrelation. In addition, due to the nature of the proposed temperature model in terms of the independence of the included processes and motivation to keep the process simple, the model assumes the independence of the HDDs. With this assumption, the approximated distribution of CHDD can be found using

$$\text{CHDD} \sim N(PB - PM - PT^*, PV). \quad (21)$$

2.3.2. Measuring HDD. HDDs are clearly contingent claims on how temperature deviates from a base temperature. As one way to find the expected value of an HDD, this study will first find its Fourier transform. Then, the inverse Fourier transform will be applied to both the HDD's Fourier transform and the characteristic function of temperature [24].

Let $x = T_t$, $w(x)$ is HDD's payoff function given in (2), Base = B, and $\widehat{w}(z) = \mathcal{F}[w(x)]$, $z \in \mathbb{C}$ is its generalized Fourier transform. Then, $\widehat{w}(z) = \int_{-\infty}^{\infty} \exp(izx)w(x)dx$.

Then,

$$\widehat{w}(z) = -\frac{e^{izB}}{z^2}, \quad \text{Im } z < 0. \quad (22)$$

Now, the inversion will be applied to $\widehat{w}(z)\theta_T(-z)$, where $\widehat{w}(z)$ is defined in (22) and θ_T is the characteristic function defined in (15). Let temperature in (9) be defined in shorthand notation as $T_t = T^* + \Lambda_t$, where T^* is defined as in (16) and $\Lambda_t = e^{-bt} \int_0^t e^{bu} dL_u$.

The characteristic function of Λ_t can be obtained from (15) and written as $\theta_\Lambda(u) = \exp(iu((1 - e^{-bt})/b)\{y_0e^{-\alpha} + \lambda_y\mu_y((1 - e^{-\alpha})/\alpha) + z_0e^{-\beta} + \lambda_z\mu_z((1 - e^{-\beta})/\beta)\} - (1/2)u^2((1 - e^{-2bt})/2b)\{C + \lambda_y\delta_Y^2((1 - e^{-2\alpha})/2\alpha) + \lambda_z\delta_z^2((1 - e^{-2\beta})/2\beta)\})$.

Then, $E[\text{HDD}] = E[(1/2\pi) \int_{iv-\infty}^{iv+\infty} e^{-izT_t} \widehat{w}(z) dz]$, where E represents expectations

$$E[\text{HDD}] = \frac{e^{-T^*}}{2\pi} \int_R e^{-iuT^*} \widehat{w}(u - i) \theta_\Lambda(-u + i) du. \quad (23)$$

However, it was not possible to evaluate this integral analytically, and thus the elliptic package of the *R* statistical software package [25] was applied to evaluate the integral numerically. The results indicated that the integral is equal to $B - M - T^*$; therefore,

$$E[\text{HDD}] = B - M - T^*. \quad (24)$$

2.3.3. Numerical Estimates. The success of the proposed temperature model and (24) were tested in terms of forecasting Cooling Degree Day (CDD), which is another index based on temperature, and HDD values for the 12 cities listed in Tables 1 and 2. CDD is calculated as $\text{CDD}_i = \max(0, T_i - \text{Base})$, where Base is a predetermined temperature level and T_i is the average temperature calculated as in (1). Cumulative CDD (CCDD) is calculated using $\sum_{i=1}^N \text{CDD}_i$, where CDD_i is calculated as in the previous sentence and N is the

TABLE 1: One-year-ahead prediction and error values for HDDs.

City	Actual	The current model	Years*	Error (%)	Campbell	Error (%)	Benth	Error (%)	HBA	Error (%)	Equation (24)	Error (%)
Ankara	6047	4980	5	-17.65	6203	2.58	5377	-11.08	5940	-1.77	6398	5.81
Beijing	5880	5705	5	-2.98	5571	-5.26	5039	-14.30	5319	-9.54	6218	5.75
Cairo	573	322	29	-43.81	746	30.20	353	-38.39	701	22.34	718	25.31
Chicago	5980	5814	15	-2.78	6098	1.97	5849	-2.19	6337	5.97	3560	-40.67
Dallas	2151	1777	34	-17.39	2202	2.37	1767	-17.85	2176	1.16	1118	-48.02
Istanbul	3607	2872	9	-20.38	3808	5.57	2972	-17.61	3485	-3.38	4520	25.31
LA	1441	1250	21	-13.26	1127	-21.79	1119	-22.35	1244	-13.67	1608	11.59
New York	4378	4391	7	0.30	4699	7.33	4516	3.15	4683	6.97	4949	13.04
Paris	3940	4988	6	26.60	5006	27.06	4190	6.35	4781	21.35	4902	24.42
Sydney	1277	1000	37	-21.70	1381	8.14	1002	-21.54	1343	5.17	1086	-14.96
Tokyo	2925	2648	8	-9.47	3238	10.70	2671	-8.68	764	-73.88	3309	13.13
Washington	3572	3758	34	5.21	3751	5.01	3724	4.26	3851	7.81	3749	4.96

*Number of years of data included in estimations.

TABLE 2: One-year-ahead prediction and error values for CDDs.

City	Actual	The current model	Years*	Error (%)	Campbell	Error (%)	Benth	Error (%)	HBA	Error (%)
Ankara	426	742	5	74.18	175	-58.92	412	-3.29	417	-2.11
Beijing	1576	1542	19	-2.16	1214	-22.97	1647	4.51	1447	-8.19
Cairo	3221	4243	5	31.73	2763	-14.22	3427	6.40	3218	-0.09
Chicago	1071	771	7	-28.01	647	-39.59	866	-19.14	909	-15.13
Dallas	3585	2556	5	-28.70	2353	-34.37	2903	-19.02	2792	-22.12
Istanbul	1428	1297	5	-9.17	792	-44.54	1224	-14.29	1089	-23.74
LA	426	490	32	15.02	569	33.57	495	16.20	681	59.86
New York	1297	1104	7	-14.88	906	-30.15	988	-23.82	1046	-19.35
Paris	264	118	8	-55.30	63	-76.14	189	-28.41	314	18.94
Sydney	1308	1251	9	-4.36	840	-35.78	1206	-7.80	1157	-11.54
Tokyo	1832	1608	18	-12.23	1282	-30.02	1716	-6.33	434	-76.31
Washington	1927	1739	8	-9.76	1451	-24.70	1503	-22.00	1609	-16.50

*Number of years of data included in estimation.

time horizon, which is generally a month or season. The test assumes a base temperature of 18 degrees Celsius and proceeds in the following manner:

Data. Temperature data is available for 12 cities covering 38 years from 1974 to 2011. The first part uses 37 years of data to estimate the parameters. The temperature data for 2011 were used to compare with one-year-ahead predictions. Temperature data were obtained from the National Climatic Data Center.

Design. A simulation was designed to run in two dimensions: the first on different cities and the second to capture changes

in the parameters through time for each city. In this respect, the initial simulations used the last 5 years of data. They were then continued by including one more year of data to the existing data in each turn where HDD and CDD estimates were calculated. A turn consisted of 10,000 simulation runs. Then, the simulation results were compared with actual HDD and CDD values. Finally, the parameters of the year that offered the best estimates of the HDDs and CDDs were chosen for use in the one-year-ahead predictions.

Discretization. Discretization was done using Euler approximation [26–28] as follows:

$$T_{t+1} = T_t + T_{t+1}^m - T_t^m + b(T_t^m - T_t) + H_{t+1} + Y_t - \alpha Y_t + (Q_{t+1} - Q_t) + Z_t - \beta Z_t + (R_{t+1} - R_t) \quad (25)$$

where $H_{t+1} = \sqrt{\gamma_0 + \gamma_1 H_t} \epsilon_t$, γ_0 and γ_1 are being ARCH parameters and $\epsilon_t \sim N(0, 1)$.

Parameter Estimation. Parameters were estimated as defined in [1, 27, 29].

Simulation of Jumps. Because they have a different structure, jumps were simulated separately and the results added to the discretized model. For this aim, the jumps were detected by first removing the mean from the actual data and selecting the values above two standard deviations. These jumps were then separated into two categories: single and multiple jumps to constitute fast and slow mean reverting jumps, respectively. The sample means and sample standard deviations of these two jump groups were found to simulate jump sizes. In addition, the intensities were found by $\hat{\lambda} = \text{no. of jumps/no. of observations}$. The simulation consisted of 10,000 runs, during which the jump times were first found by using intensities. Then, for each run, random draws were taken using means and standard deviations obtained from data. Finally, the discretized jumps were added to the discretized jump model.

After these phases, one-year-ahead predictions were conducted along with three other models: the Campbell and Diebold [4] model (Campbell Model), the Benth and Šaltytė-Benth [2] model (Benth Model), and the Historical Burn Analysis (HBA) model that calculates historical averages. In addition, HDD predictions were calculated based on (24). This yielded the results in Tables 1 and 2. In the Appendix, the parameter estimates for HDD and CDD values are presented, in Tables 3 and 4, respectively, including the time period used to obtain these parameters. In addition, statistical test values of these parameters are shown in Tables 5 and 6.

2.3.4. Analysis of Numerical Estimates. The best estimates of HDDs and CDDs were obtained for different periods as shown in the Appendix. This is mainly a characteristic of the temperature since it changes in its long-term behavior. However, this may not be a good way to use all existing data for a city. Instead, every location must be scanned and evaluated for different time periods to obtain the best

prediction results. In addition, the current model is equally successful in HDD and CDD predictions.

Having a good estimate of HDD and CDD values does not necessarily correspond to the best fit of the model to temperature data. This may be because including the jumps may result in a better estimation of index values while deteriorating the fit of the model to the data.

The current model demonstrates its capacity related to the changing conditions of the temperature data. For example, in Chicago, both jump types were statistically significant and the model predicted HDDs accurately. On the other hand, Tokyo did not have any jumps during entire period, and the current model was still able to make accurate predictions for HDDs. Interestingly, the Historical Burn Analysis containing long-term HDD and CDD averages provided successful predictions. These results suggest that temperatures do not change significantly for certain locations.

As expected, the approximated HDD calculations obtained from (24) were less accurate than the simulations. Nevertheless, the predictions based on the equation were still successful. The estimated HDDs of Los Angeles and Washington were better than any other model. Finally, while this study conducted comparisons for only 12 stations, there are 125,000 weather stations worldwide. It is therefore impossible to say that one model is superior to the others, though the current model and (24) were successful in certain locations and periods, which merits evaluation.

3. Pricing

This section addresses temperature risk and its differences from classical asset risks before providing a fair price for a temperature-based derivative written of HDD with real probabilities. Next, this section shows that the price of the derivative will be super-hedging when using risk-neutral probabilities. Finally, a personal price will be developed based on personal temperature risk.

TABLE 3: Parameter estimation for HDD calculations.

City	Year*	A	B	C	Phase	b	Gamma_0	Gamma_1	Fast_Inten	Fast_Mean	Fast_sd**	Slow_Inten	Slow_Mean	Slow_sd	Beta
Ankara	5	50.19	0.001854	21.32	-2.77	0.193962	95.952994	0.9637143	0.0016438	-0.5543834	0.4924429	0.0054795	-6.1314104	5.1345993	0.3835657
Beijing	5	56.73248	-0.00176	27.72651	-2.89336	0.372348	66.177766	0.9800345	0	0	0	0.00109589	-3.161959	2.420366	0.47568
Cairo	29	70.23189	0.000344	13.68758	-2.72869	0.345847	239.81369	0.9547672	0.0042513	-0.4124696	1.585721	0.00302315	-1.690717	2.075651	0.703044
Chicago	15	50.57	0.00014	24.71	-2.79	0.271967	117.25	0.961314	0.004414	-3.119119	2.524564	0.005632	-6.375580	5.517740	0.343206
Dallas	34	65.31678	0.000232	20.00168	-2.83072	0.298336	294.61489	0.936274	0.0045931	-2.4987494	2.391683	0.00604351	-6.1932845	4.966445	0.291944
Istanbul	9	58.54	0.00081	18.12	-2.67	0.237101	184.55	0.947974	0.004262	-0.958549	2.047743	0.002740	-3.133173	1.813471	0.562597
Los Angeles	21	63.82858	-0.0001	6.651395	-2.4675	0.230077	338.75801	0.9166473	0.0062622	0.3575345	1.352837	0.01108937	0.81062689	2.798516	0.460403
New York	7	54.63592	0.000548	21.77908	-2.71524	0.286476	165.02405	0.9447895	0.0015656	-1.6825421	2.360227	0.00508806	-3.8112152	3.30503	0.684576
Paris	6	54.58335	-0.00122	15.45061	-2.83361	0.21059	147.12152	0.9490173	0.0054795	-0.0954452	1.895953	0.00547945	-1.67898	3.505224	0.554255
Sydney	37	63.451	0.000154	9.494916	-2.8139	0.647759	526.99911	0.876059	0.0146612	1.5064313	2.606529	0.0033321	1.5601199	3.226385	0.911471
Tokyo	8	61.22	0.000375	18.38	-2.61866	0.530895	180.28	0.954846	0	0	0	0	0	0	0
Washington	34	58.62177	3.04E - 05	21.75468	-2.79459	0.279335	173.30119	0.9524289	0.0045125	-1.9805768	1.755655	0.00354553	-4.9014241	4.06658	0.388133

*Year represents amount of data in years that produced best estimation results.

** sd means standard deviation.

TABLE 4: Parameter estimation for CDD calculations.

City	Year*	A	B	C	Phase	b	Gamma_0	Gamma_1	Fast_Inten	Fast_Mean	Fast_sd**	Slow_Inten	Slow_Mean	Slow_sd	Beta
Ankara	5	50.19	0.001854	21.32	-2.77	0.193962	95.952994	0.9637143	0.0016438	-0.5543834	0.4924429	0.0054795	-6.1314104	5.1345993	0.3835657
Beijing	19	54.96688	3.03E - 06	2714738	-2.91682	0.351405	63.312073	0.980961	0.000721	-1.5276715	1.531992	0.00100937	-2.9315846	2.019894	0.385117
Cairo	5	71.94036	0.002198	13.39397	-2.74021	0.36773	104.83607	0.983445	0.0021918	0.3088935	1.644226	0.00273973	-1.7313942	0.945926	0.544186
Chicago	7	51.70	-0.00034	25.18	-2.80	0.253619	108.27	0.963663	0.003131	-3.281461	3.025694	0.005871	-5.240072	4.234712	0.397902
Dallas	5	69.85579	-0.00172	20.02312	-2.83355	0.269534	314.92838	0.934	0.0032877	-3.2390974	1.069774	0.00657534	-3.3012769	3.345035	0.4432
Istanbul	5	59.32	0.00134	18.10	-2.69	0.264510	183.29	0.949545	0.002192	0.067349	0.601966	0.002740	-2.406968	1.833246	0.561837
Los Angeles	32	63.81324	-5.64E - 05	6.699702	-2.46624	0.208063	312.82723	0.9232099	0.0058219	0.3913695	1.425177	0.00984589	1.02398907	3.358289	0.40139
New York	7	54.63592	0.000548	21.77908	-2.71524	0.286476	165.02405	0.9447895	0.0015656	-1.6825421	2.360227	0.00508806	-3.8112152	3.30503	0.684576
Paris	8	54.66317	-0.00085	15.5993	-2.83742	0.218952	150.30161	0.9485509	0.005137	-0.4207048	1.581089	0.00513699	-0.1355112	4.770146	0.255761
Sydney	9	65.07105	0.000171	9.53934	-2.86309	0.535114	426.96927	0.9023923	0.0152207	1.6436984	1.81751	0.00182648	2.58516471	2.615411	0.764766
Tokyo	18	61.60	3.13E - 05	18.53	-2.63	0.497182	214.39	0.944722	0	0	0	0	0	0	0
Washington	8	57.7754	0.000801	22.01289	-2.78212	0.284276	169.0818	0.9526032	0.0041096	-1.6365612	1.493845	0.00376712	-3.2540612	2.309414	0.561109

*Year represents amount of data in years that produced best estimation results.

** sd means standard deviation.

TABLE 5: *P* values of parameters for HDD calculations.

City	A	B	C	Phase	Gamma_0	Gamma_1	Fast_Mean	Slow_Mean
Ankara	0	0	0	0	NS	0.00263	0.1609	0
Beijing	0	0	0	0	NS	0.0082	0.3174	0.02142
Cairo	0	0	0	0	NS	NS	0.0879	0
Chicago	0	NS	0	0	0	0	0	0
Dallas	0	0	0	0	NS	0	0	0
Istanbul	0	0	0	0	NS	NS	0.1021	0
Los Angeles	0	0	0	0	NS	NS	0.07382	0
New York	0	<0.001	0	0	NS	<0.001	0.2038	0
Paris	0	0	0	0	NS	<0.01	0.8557	0.01075
Sydney	0	0	0	0	NS	<0.05	0	0
Tokyo	0	0	0	0	NS	NS	NA	NA
Washington	0	<0.05	0	0	NS	0	0	0

NS: not statistically significant.

NA: not available.

TABLE 6: *P* values of parameters for CDD calculations.

City	A	B	C	Phase	Gamma_0	Gamma_1	Fast_Mean	Slow_Mean
Ankara	0	0	0	0	NS	0.00263	0.1609	0
Antalya	0	0	0	0	NS	NS	0.3143	0.7321
Beijing	0	NS	0	0	NS	0	0.096	0
Bursa	0	0	0	0	NS	NS	0.3216	0
Cairo	0	0	0	0	NS	NS	0.6717	0.001395
Chicago	0	NS	0	0	<0.01	0	0.03218	0
Dallas	0	0	0	0	NS	NS	0.01897	0
Istanbul	0	0	0	0	NS	NS	0.7978	0
Los Angeles	0	0	0	0	NS	NS	0.02789	0
New York	0	<0.001	0	0	NS	<0.001	0.2038	0
Paris	0	0	0	0	NS	<0.01	0.3038	0.8347
Sydney	0	<0.01	0	0	NS	NS	0	0.0056
Tokyo	0	<0.01	0	0	NS	<0.01	NA	NA
Washington	0	0	0	0	NS	<0.001	0.009	0

NS: not statistically significant.

NA: not available.

3.1. Temperature Risk. In simple terms, temperature risk is volume risk that affects sales [13]. This study will first focus on a single company with an obvious exposure to temperature to reveal the relationship between temperature index and the company's sales. To maintain the focus on the relationship between sales and temperature, a very simple linear model that omits other possible factors that may affect sales will be introduced. As a candidate company, consider a retail gas seller concerned about sales and profit in the following January. The company's risk will be measured through the effect of sales on profits. Consider the following:

$$ES = a + bCHDD \quad \text{where ES is expected sales.} \quad (26)$$

A positive value for a in (26) indicates that the company can sell a certain amount of its products, even in the case of zero CHDD. Similarly, a positive value for b in the same equation represents a positive relationship between sales and CHDDs.

The relationship between CHDDs and profit can be established using cost and revenues.

$$\text{Expected Cost} = C = \theta + \text{Expected Sales} \quad (27)$$

where θ is a constant

$$\text{Expected Revenue} = R = \text{Price} * \text{Expected Sales} \quad (28)$$

where Price represents the price of the product sold

$$\begin{aligned} \text{Expected Profit} = P &= R - C = \text{Price} * ES - \theta - ES \\ &= ES(\text{Price} - 1) - \theta. \end{aligned} \quad (29)$$

To guarantee a positive profit after a certain amount of CHDDs is realized, assume that $\text{Price} > 1$ and constant.

The aim is to construct a relationship between CHDD and profit functions to ascertain the magnitude of the effect of temperature on the company. Figure 1 illustrates this

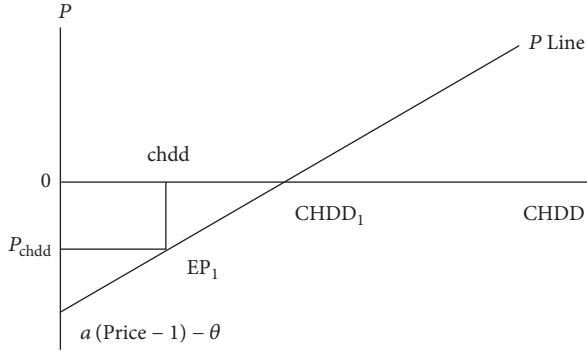


FIGURE 1: Relationship between expected profit and CHDD.

relationship, with the assumption that $\theta > a * \text{Price}$ initially provides negative profits.

Figure 1 constructs a relationship between CHDD and profit such that every CHDD value now represents a monetary value in terms of positive and negative profits. Figure 1 shows a deterministic relationship between CHDD level and profit. The CHDD value for the proposed period is unknown.

Now, the company's risk can be defined as having a low value of CHDD for a certain period that leads to a loss. In other words, the company's risk will fall left of CHDD_1 and will be covered by the P Line. This possible loss will be referred to as temperature risk (TR) and is equal to the area covered by 0 $\text{CHDD}_1 a(\text{Price} - 1) - \theta$ in Figure 1. This is actually the total TR (TTR) and can be realized if CHDD for the next January becomes zero. In reality, the CHDD for the next January is not known because it is a stochastic value. Therefore, TR can be written as

$$\text{TR} = \text{TTR} - \int_0^{\text{chdd}} P d\text{CHDD} = \int_{\text{chdd}}^{\text{CHDD}_1} P d\text{CHDD}, \quad (30)$$

where chdd is unknown value of CHDD for a certain period.

It is clear that the magnitude of TR will depend on business type and size. For example, for a gas company, a decrease in the index value means a lower value in sales. However, at the same time, this decrease may result in an increase in the sales of a beverage company. The magnitude of the decrease or the increase in profits, on the other hand, will be directly related to the size of the business.

3.2. An Approximated Fair Price of a Temperature Based Put Option. Under linear approximation, the temperature has a normal distribution with mean M^* and variance V , as defined in (18) and (19), respectively. Consider the following:

$$\begin{aligned} \text{CHDD} = & \text{HDD of Day 1} + \text{HDD of Day 2} + \dots \\ & + \text{HDD of Day } P. \end{aligned} \quad (31)$$

The distribution function is given in (21). Let K be the strike value, $x = \text{CHDD}$, let $f(x)$ be the probability density function of CHDD, and tick the monetary value for each

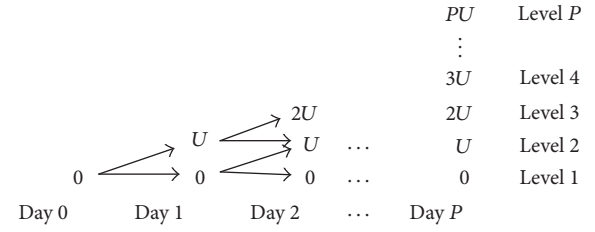


FIGURE 2: Construction of a CHDD tree using a binomial model.

CHDD. Then, the value of a put option on CHDD will be equal to

$$\begin{aligned} E[\max(K - x, 0)] &= \int_{-\infty}^K (K - x) f(x) dx \\ E[\max(K - x, 0)] &= e^{-r(T-t)} \left[(K - PB + PM^*) N(K - PB + PM^*) \right. \\ &\quad \left. + \frac{\sqrt{PV}}{\sqrt{2\pi}} e^{-A^2/2PV} \right] * \text{tick}. \end{aligned} \quad (32)$$

Equation (32) provides an actuarial price since it is based on expected values. Derivative pricing using real probabilities may require an evaluation of risk premiums. The literature offers only inconclusive discussions. For example, Hull [19] states that it is possible to calculate the payoffs of weather derivatives with real probabilities because these derivatives have no systemic risk. Turvey [30] supports this idea. On the other hand, Chincarini [31] examines the efficiency of weather futures in CME in HDD and CDD futures assuming an efficient market and risk premiums varying from negative to positive values across cities. In addition, Cao and Wei [13] highlighted the importance of the market price of risk for weather derivatives. This study will follow a different path by evaluating the trading behavior of a candidate company instead of the company's risk behavior. The next section will show the outcome of using risk-neutral probabilities in (32).

3.3. Risk-Neutral Pricing. In this part, the put option of the previous section will be priced using risk-neutral probabilities in a binomial model. CHDD is clearly composed of HDDs. A closer look at HDDs reveals that realizations of HDDs can be represented by a binomial model such that if HDD is realized, there will be an upward movement as in the binomial model; otherwise, in case of a downward movement, there will be 0. The calculations were done according to Björk [32]. For the up movement, the best candidate for HDD will be the mean value of an HDD calculated using (24). Let $U = (24)$. Figure 2 represents the binomial model.

The probability of an upward movement, p_u , is found by $\int_{-\infty}^B f(x) dx$, while the probability of downward movement is equal to $p_d = 1 - p_u$. Let risk-free rate $R = 0$. The model satisfies the condition of being arbitrage-free in the form of $d \leq 1 + R \leq u$ by definition. The martingale measure for the

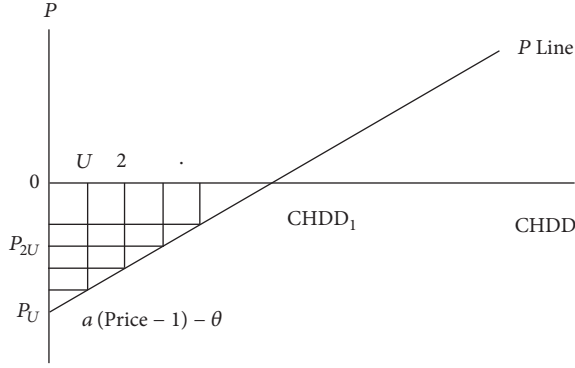


FIGURE 3: Evolution of TR.

current model is $CHDD_t = E^Q[CHDD_{t+1}]$. As in Björk [32], if it is set $K = PU$, the value of the option will be equal to K itself.

The above results from the fact that CHDD is a summation process and certainly a submartingale compared to the underlying asset of an ordinary option. The only possibility to obtain a martingale form of the underlying process CHDD is then to consider the whole summation process and reflect it as a constant. Reexamining Figure 1 reveals one interesting implication of this result. Figure 3 is a combination of Figures 1 and 2.

In Figure 3, each level of TR is connected to upward movements created by HDDs. This time, TR becomes a super-martingale; when TR is converted into a martingale, it will be equal to TTR, a situation known as super-hedging. Put simply, it means hedging the total risk. In pricing, the price of the hedge will be equal to total risk. Using the total risk instead of actual risk will definitely prevent any form of transaction, for both hedge suppliers and demanders. Therefore, using risk-neutral probabilities is not ideal for pricing a temperature-based derivative. With these unclear results from the risk premiums and the inappropriateness of using risk-neutral probabilities, the current study offers the following for pricing temperature-based derivatives.

3.4. A New Setup for Pricing. In this setup, any economic entity, an individual or a company, will try to achieve an objective. In this case, the natural objective for the hypothetical company is to maximize its expected profit given in (29). Then, the issue is determining the conditions under which this company will maximize its profits. The answer will reveal the company's trading behavior. Consider Proposition 1.

Proposition 1. *The company will buy a put option if the following condition is satisfied:*

$$E[\max(K - CHDD, 0)] N(K - PB + PM^*) \geq C. \quad (33)$$

$E[\max(K - CHDD, 0)]$ is given in (32). $N(K - PB + PM^*)$, which is the probability of being in the money, is also obtained from (32). C is the cost of the put option. The tick value is equal to \$1.

Proof. Equation (29) is the expected profit when there is no trade for options. If there is a chance to trade an option, the company will prefer a put option with a strike value equal to $CHDD_1$ from Figure 1. Assume that the company buys a put option equal to ϵ with a cost of C . There will be two states at the end of the determined period depending on the payout of the option. The setup is given in the following.

Let x represent CHDD. Then, one has the following.

State 1

$$\begin{aligned} & ((a + bE[x])(\text{Price} - 1) - \theta - \epsilon C \\ & + \epsilon E[\max(K - x, 0)]) \end{aligned} \quad (34)$$

where E represents expectation.

State 2

$$\begin{aligned} & ((a + bE[x])(\text{Price} - 1) - \theta - \epsilon C) \\ & \text{where } E \text{ represents expectation.} \end{aligned} \quad (35)$$

The probability $N(K - PB + PM^*)$ also defines the probability of State 1. Then, the probability of State 2 will be $1 - N(K - PB + PM^*)$. Under these probabilities, the expected value of the two states will be

$$\begin{aligned} & ((a + bE[x])(\text{Price} - 1) - \theta - \epsilon C \\ & + \epsilon E[\max(K - x, 0)]) N(K - PB + PM^*) \\ & + ((a + bE[x])(\text{Price} - 1) - \theta - \epsilon C) (1 \\ & - N(K - PB + PM^*)). \end{aligned} \quad (36)$$

The company will enter a trade for a put option if (36) is greater than or equal to the no trade case such that

$$\begin{aligned} & ((a + bE[x])(\text{Price} - 1) - \theta - \epsilon C \\ & + \epsilon E[\max(K - x, 0)]) N(K - PB + PM^*) \\ & + ((a + bE[x])(\text{Price} - 1) - \theta - \epsilon C) (1 \\ & - N(K - PB + PM^*)) \geq (a + bE[x])(\text{Price} - 1) \\ & - \theta. \end{aligned} \quad (37)$$

Additionally, the profit function in the no trade case is also reflected in the left hand side of the equation with a probability of 1. Consequently, subtracting the right hand side from the left hand side will result in

$$-\epsilon C + \epsilon E[\max(K - x, 0)] N(K - PB + PM^*) \geq 0. \quad (38)$$

□

Although the above calculations are given for a candidate company, since profit function dropped out and the tick value is equal to \$1, (33) defines a general case valid for any company dealing with a put option with $K = CHDD_1$. Thus, the following definition is given for the case $E[\max(K - x, 0)] N(K - PB + PM^*) = C$.

Definition 2. When $E[\max(K - x, 0)]N(K - PB + PM^*) = C$, C is the general price, which is valid for any economic entity.

The above setup can be extended by the concept of a shadow price, which describes the effects of the resources in a production process on profit. In the current case, the profit function is a deterministic function that is a payoff function of the index value CHDD. Therefore, CHDD can be seen as the resource that produces the profit. It is possible to measure the effect of one unit of change in CHDD on profit and use it as the price of one unit of CHDD. Again, assuming $CHDD = x$, consider the following:

$$\frac{d\text{Profit}(x)}{dx} = b(\text{Price} - 1). \quad (39)$$

The value given in (39) is a good candidate for the tick value mentioned in (32). Now, using (32) and (39), a new definition can be given.

Definition 3. The personalized price for the put option $E[\max(K - x, 0)]$ is given by

$$E[\max(K - x, 0)]N(K - PB + PM^*)(b(\text{Price} - 1)) = C. \quad (40)$$

3.5. Discussion of the New Pricing Setup. Equation (33) defines the profitable conditions for the company, of which there are three. When $E[\max(K - x, 0)]N(K - PB + PM^*) \geq C$, the company will buy the option. If $E[\max(K - x, 0)]N(K - PB + PM^*) = C$, the company will be indifferent between buying the option or doing nothing. Finally, when $E[\max(K - x, 0)]N(K - PB + PM^*) \leq C$, it is the best interest of the company to sell the option because it maximizes profit. The value of C when $E[\max(K - x, 0)]N(K - PB + PM^*) = C$ can be defined as the fair price since it does not result in a positive profit.

From here, a connection between the current approach and the utility approach can be established. Since the current setup is based on profit maximization, it coincides with the utility approach based on wealth maximization. The gain is that this statement is true for any utility function choices.

A numerical one-day-ahead estimate of temperature for the price of an HDD for Ankara was developed using (33) and (32). The mean and standard deviation of the approximated distribution were calculated according to (18) and (19). The value of C in (19) was approximated by conditional variance of the ARCH model. The tick value was taken as \$1, and the strike value K was taken as an interval from 65 to 100. The estimated values are shown in Figure 4.

As mentioned earlier, (33) defines a general price and trading behaviors for any company since the equation does not include the profit function. This generality does not give much insight into what a put option with a strike value of K actually means for a specific company. This deficiency was corrected by replacing the tick value with (40) because, unlike ordinary assets, temperature affects economic entities on different scales. Thus, a personalized price must apply for each economic entity. Moreover, (40) and Definition 3

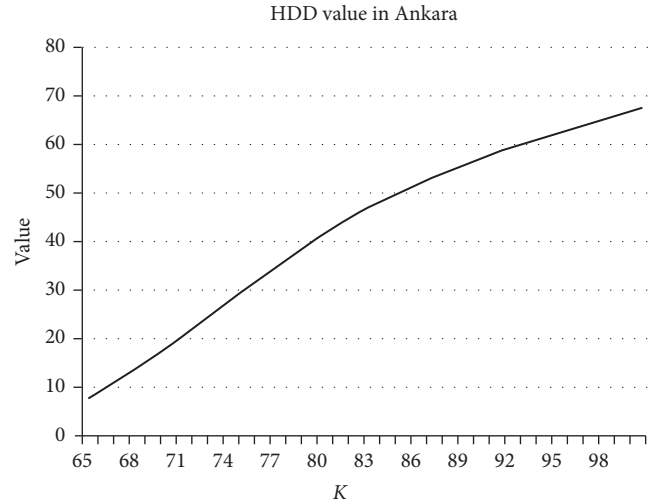


FIGURE 4: Estimated value of an HDD for Ankara.

state that the hypothetical company will enter a trade for the option if there is a possibility for arbitrage. If the fair price is available, the company will be indifferent to entering a trade or doing nothing. In the new pricing setup, the expected profit will always be the maximum, as will be the company's utility. Risk aversion will negatively affect the maximum profit and utility. Therefore, having a maximum profit and the resulting maximum utility will direct the company to follow the presented approach rather than the one of suboptimal risk aversion.

4. Conclusions

Derivatives written on temperature are based on index values obtained from temperature data, which are essentially measured as deviations of temperature from a threshold value. This makes measuring deviations from a base temperature in the form of jumps important for any temperature model for some locations. This study proposed and demonstrated a temperature model that included different kinds of jumps that was then handled using different techniques. In addition, unlike existing models that consider temperature risk as the result of the temperature itself, like in stocks, the proposed model shows that financial risk caused by temperature differs from classical asset risk, though this risk depends on the type of business. This study demonstrated a method to measure this temperature risk. Moreover, almost all of the existing pricing methods are based on risk-neutral valuations. The results from this study showed that risk-neutral valuation in temperature-based derivatives ends with super-hedging. Finally, the current study offers a pricing scheme that differs from classical pricing approaches that are based on risk-neutrality or risk-aversion concepts. Instead of utility functions, this study employs a more realistic and practical approach in terms of objective functions set by the firm itself. In return, the model provides a personalized price based on company-specific temperature risk to realize an objective in terms of profit.

Appendix

See Tables 3, 4, 5, and 6.

Disclosure

This paper is based on the Ph.D. thesis by Birhan Taştan, titled “Modeling Temperature and Pricing Weather Derivatives Based on Temperature,” which was presented at the Middle East Technical University, Institute of Applied Mathematics, in the year 2016.

Conflicts of Interest

The authors declare that there are no conflicts of interest regarding the publication of this paper.

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